#### Review of Probability Theory

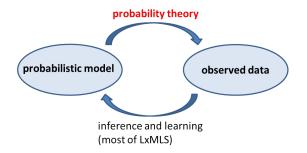
#### Mário A. T. Figueiredo

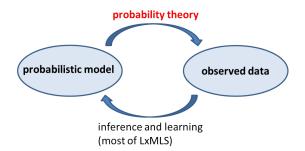
Instituto Superior Técnico & Instituto de Telecomunicações

Lisboa, Portugal

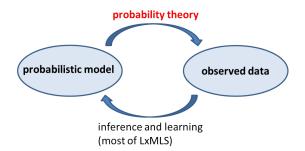
LxMLS: Lisbon Machine Learning School

July 24, 2013

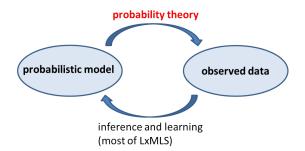




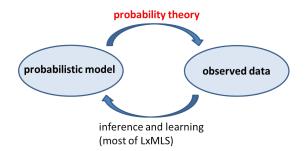
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- Great names of science: Cardano, Fermat, Pascal, Laplace, Kolmogorov, Bernoulli, Poisson, Cauchy, Boltzman, de Finetti, ...
- Natural tool to model uncertainty, information, knowledge, belief, ...
- ...thus also learning, decision making, inference, ...

### What is probability?

• Classical definition: 
$$\mathbb{P}(A) = \frac{N_A}{N}$$

...with N mutually exclusive equally likely outcomes,  $N_A$  of which result in the occurrence of A. Laplace, 1814

Example:  $\mathbb{P}(\text{randomly drawn card is }) = 13/52.$ 

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• Subjective probability:  $\mathbb{P}(A)$  is a degree of belief. *de Finetti, 1930s* 

...gives meaning to  $\mathbb{P}($  "tomorrow will rain" ).

#### Key concepts: Sample space and events

- Sample space  $\mathcal{X} =$  set of possible outcomes of a random experiment. Examples:
  - Tossing two coins:  $\mathcal{X} = \{HH, TH, HT, TT\}$
  - Roulette:  $X = \{1, 2, ..., 36\}$
  - ▶ Draw a card from a shuffled deck:  $\mathcal{X} = \{A, 2, ..., Q \diamondsuit, K \diamondsuit\}$ .

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  - ▶ Draw a card from a shuffled deck:  $\mathcal{X} = \{A\clubsuit, 2\clubsuit, ..., Q\diamondsuit, K\diamondsuit\}$ .
- An event is a subset of  $\mathcal X$

Examples:

- "exactly one H in 2-coin toss":  $A = \{TH, HT\} \subset \{HH, TH, HT, TT\}$ .
- "odd number in the roulette":  $B = \{1, 3, ..., 35\} \subset \{1, 2, ..., 36\}$ .
- "drawn a  $\heartsuit$  card":  $C = \{A\heartsuit, 2\heartsuit, ..., K\heartsuit\} \subset \{A\clubsuit, ..., K\diamondsuit\}$

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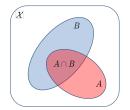
• 
$$\mathbb{P}(\mathcal{X}) = 1$$

• If  $A_1, A_2 ... \subseteq \mathcal{X}$  are disjoint events, then  $\mathbb{P}\left(\bigcup_i A_i\right) = \sum_i \mathbb{P}(A_i)$ 

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- ▶ If  $A_1, A_2 ... \subseteq \mathcal{X}$  are disjoint events, then  $\mathbb{P}(\bigcup A_i) = \sum_i \mathbb{P}(A_i)$
- From these axioms, many results can be derived. Examples:
- $\mathbb{P}(\emptyset) = 0$
- $C \subset D \Rightarrow \mathbb{P}(C) \leq \mathbb{P}(D)$
- $\blacktriangleright \mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) \mathbb{P}(A \cap B)$
- $\mathbb{P}(A \cup B) \leq \mathbb{P}(A) + \mathbb{P}(B)$  (union bound)

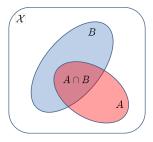


...satisfies all of Kolmogorov's axioms:

• For any 
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,  $\mathbb{P}(A|B) \ge 0$ 

•  $\mathbb{P}(\mathcal{X}|B) = 1$ 

• If 
$$A_1, A_2, ... \subseteq \mathcal{X}$$
 are disjoint, then  
 $\mathbb{P}\left(\bigcup_i A_i \middle| B\right) = \sum_i \mathbb{P}(A_i | B)$ 

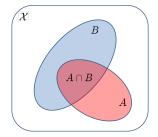


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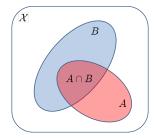
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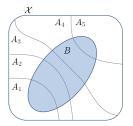
- Events A, B are independent  $(A \perp B) \Leftrightarrow \mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B)$ .
- Relationship with conditional probabilities:

$$A \perp\!\!\!\perp B \Leftrightarrow \mathbb{P}(A|B) = \mathbb{P}(A)$$

#### **Bayes Theorem**

• Law of total probability: if  $A_1, ..., A_n$  are a partition of  $\mathcal{X}$ 

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$$= \sum_{i} \mathbb{P}(B \cap A_{i})$$

$$\mathcal{X}$$

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$$A_{2}$$

$$A_{3}$$

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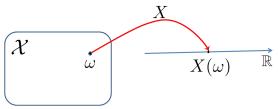
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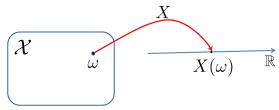
• Bayes' theorem: if  $A_1, ..., A_n$  are a partition of  $\mathcal{X}$ 

$$\mathbb{P}(A_i|B) = \frac{\mathbb{P}(B \cap A_i)}{\mathbb{P}(B)} = \frac{\mathbb{P}(B|A_i) \mathbb{P}(A_i)}{\sum_i \mathbb{P}(B|A_i) \mathbb{P}(A_i)}$$

• A (real) random variable (RV) is a function:  $X : \mathcal{X} \to \mathbb{R}$ 

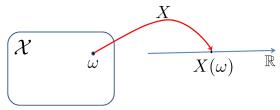


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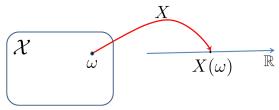
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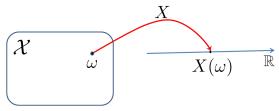


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• Example: number of head in tossing two coins,  

$$\mathcal{X} = \{HH, HT, TH, TT\},$$
  
 $X(HH) = 2, X(HT) = X(TH) = 1, X(TT) = 0.$   
Range of  $X = \{0, 1, 2\}.$ 

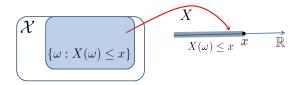
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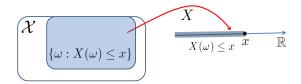


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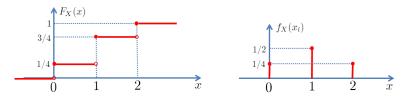
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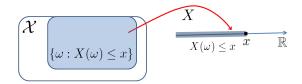
• Example: distance traveled by a tossed coin; range of  $X = \mathbb{R}_+$ .



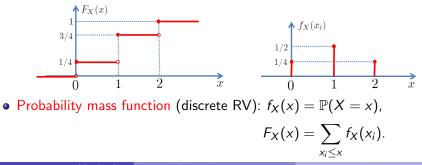


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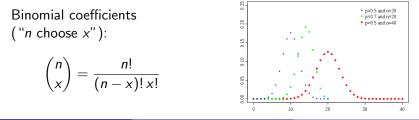
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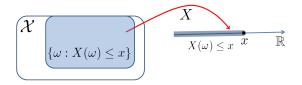
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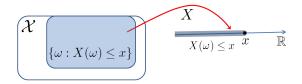
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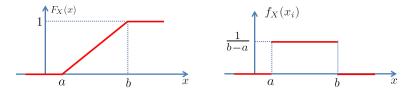
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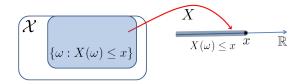
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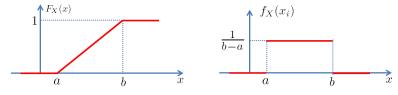
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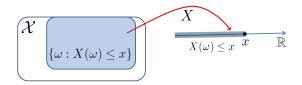


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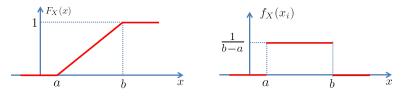


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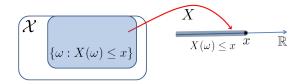
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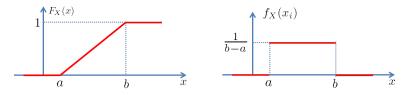
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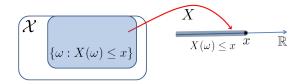
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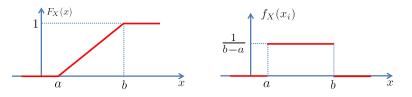
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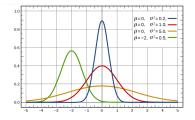
Important Continuous Random Variables

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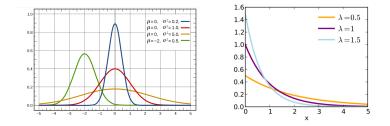
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• Exponential:  $f_X(x) = \operatorname{Exp}(x; \lambda) = \begin{cases} \lambda e^{-\lambda x} & \Leftarrow x \ge 0 \\ 0 & \Leftarrow x < 0 \end{cases}$ 

• Expectation: 
$$\mathbb{E}(X) = \begin{cases} \sum_{i} x_i f_X(x_i) & X \in \{x_1, ..., x_K\} \subset \mathbb{R} \\ \int_{-\infty}^{\infty} x f_X(x) dx & X \text{ continuous} \end{cases}$$

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, for  $x \in \{0, 1\}$ .  
 $\mathbb{E}(X) = 0 (1-p) + 1 p = p$ .

• Example: Binomial, 
$$f_X(x) = \binom{n}{x} p^x (1-p)^{n-x}$$
, for  $x \in \{0, ..., n\}$ .  
 $\mathbb{E}(X) = n p$ .

• Example: Gaussian,  $f_X(x) = \mathcal{N}(x; \mu, \sigma^2)$ .  $\mathbb{E}(X) = \mu$ .

• Linearity of expectation:  $\mathbb{E}(X + Y) = \mathbb{E}(X) + \mathbb{E}(Y); \quad \mathbb{E}(\alpha X) = \alpha \mathbb{E}(X), \ \alpha \in \mathbb{R}$ 

Mário A. T. Figueiredo (IST & IT)

LxMLS 2013: Probability Theory

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$$\mathbb{E}(g(X)) = \begin{cases} \sum_{i} g(x_i) f_X(x_i) & X \text{ discrete, } g(x_i) \in \mathbb{R} \\ \int_{-\infty}^{\infty} g(x) f_X(x) dx & X \text{ continuous} \end{cases}$$

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• Probability as expectation of indicator,  $\mathbf{1}_A(x) = \begin{cases} 1 & \Leftarrow x \in A \\ 0 & \Leftarrow x \notin A \end{cases}$ 

$$\mathbb{P}(X \in A) = \int_A f_X(x) \, dx = \int \mathbf{1}_A(x) \, f_X(x) \, dx = \mathbb{E}(\mathbf{1}_A(X))$$

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• Also valid in the mixed case (e.g., X continuous, Y discrete).

# Joint, Marginal, and Conditional Probabilities: An Example

• A pair of binary variables  $X, Y \in \{0, 1\}$ , with joint pmf:

$f_{X,Y}(x,y)$	Y = 0	Y = I		
X = 0	1/5	2/5		
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• Conditional probabilities:

$f_{X Y}(x y)$	Y = 0	Y = I	$f_{Y X}(y x)$	Y = 0	Y = I
X = 0	2/3	4/7	X = 0	1/3	2/3
X = I	1/3	3/7	X = 1	1/4	3/4

#### An Important Multivariate RV: Multinomial

• Multinomial:  $X = (X_1, ..., X_K)$ ,  $X_i \in \{0, ..., n\}$ , such that  $\sum_i X_i = n$ ,

$$f_X(x_1,...,x_K) = \begin{cases} \binom{n}{x_1 x_2 \cdots x_K} p_1^{x_1} p_2^{x_2} \cdots p_k^{x_K} & \Leftarrow \sum_i x_i = n \\ 0 & \Leftarrow \sum_i x_i \neq n \end{cases}$$

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- Generalizes the binomial from binary to K-classes.
- Example: tossing *n* independent fair dice,  $p_1 = \cdots = p_6 = 1/6$ .  $x_i =$  number of outcomes with *i* dots. Of course,  $\sum_i x_i = n$ .

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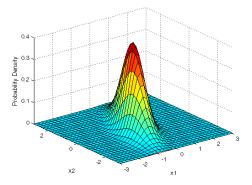
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$$\operatorname{cov}(X,Y) = \mathbb{E}\Big[ (X - \mathbb{E}(X)) (Y - \mathbb{E}(Y)) \Big] = \mathbb{E}(X Y) - \mathbb{E}(X) \mathbb{E}(Y)$$

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- In the Bayesian philosophy, all the knowledge about X is carried by

$$f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x) f_X(x)}{f_Y(y)} = \frac{f_{Y,X}(y,x)}{f_Y(y)}$$

...the posterior (or a posteriori) pdf/pmf.

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- The optimal Bayesian decision minimizes the expected loss:

$$\widehat{x}_{\mathsf{Bayes}} = \arg\min_{\widehat{x}} \mathbb{E}[L(\widehat{x}, X) | Y = y]$$

where

$$\mathbb{E}[L(\widehat{x}, X)|Y = y] = \begin{cases} \int L(\widehat{x}, x) f_{X|Y}(x|y) dx, & \text{continuous (estimation)} \\ \sum_{x} L(\widehat{x}, x) f_{X|Y}(x|y), & \text{discrete (classification)} \end{cases}$$

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- Optimal Bayesian decision:

$$\begin{split} \widehat{x}_{\mathsf{Bayes}} &= \arg\min_{\widehat{x}} \sum_{x=1}^{K} \mathcal{L}(\widehat{x}, x) \, f_{X|Y}(x|y) \\ &= \arg\min_{\widehat{x}} \left( 1 - f_{X|Y}(\widehat{x}|y) \right) \\ &= \arg\max_{\widehat{x}} f_{X|Y}(\widehat{x}|y) \; \equiv \; \widehat{x}_{\mathsf{MAP}} \end{split}$$

MAP = maximum a posteriori criterion.

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MAP = maximum a posteriori criterion.

 Same criterion can be derived for continuous X, using lim<sub>ε→0</sub> L<sub>ε</sub>(x̂, x), where L<sub>ε</sub>(x̂, x) = 0, if |x̂ − x| < ε, and 1 otherwise.</li>

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- From Bayes law:

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...only need to know posterior  $f_{X|Y}(x|y)$  up to a normalization factor.

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- Also common to write:  $\hat{x}_{MAP} = \arg \max_{x} \log f_{Y|X}(y|x) + \log f_{X}(x)$
- If the prior if flat,  $f_X(x) = C$ , then,

$$\widehat{x}_{\mathsf{MAP}} = rg\max_{x} f_{Y|X}(y|x) \equiv \widehat{x}_{\mathsf{ML}}$$

ML = maximum likelihood criterion.

• Observed *n* i.i.d. (independent identically distributed) Bernoulli RVs:  $Y = (Y_1, ..., Y_n)$ , with  $Y_i \in \{0, 1\}$ .

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- Example: n = 10, observed y = (1, 1, 1, 0, 1, 0, 0, 1, 1, 1),  $\hat{x}_{ML} = 7/10$ .

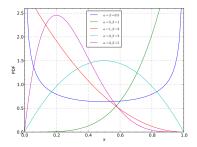
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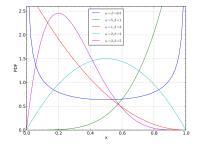
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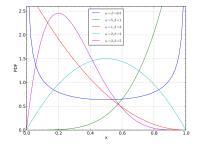
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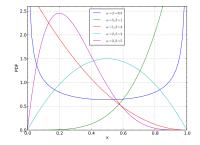
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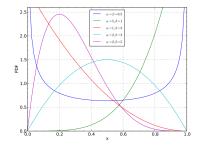
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• Example: 
$$\alpha = 4$$
,  $\beta = 4$ ,  $n = 10$ ,  $y = (1, 1, 1, 0, 1, 0, 0, 1, 1, 1)$ ,

$$\widehat{x}_{MAP} = 0.625 \text{ (recall } \widehat{x}_{ML} = 0.7 \text{)}$$



## Another Classical Statistical Inference Criterion

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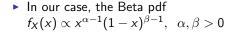
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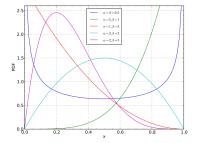
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• Does not apply to classification problems.

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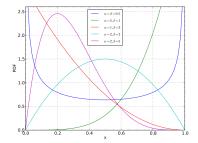
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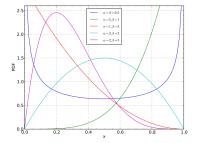
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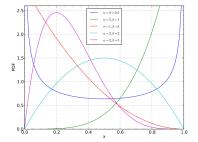
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• Conjugate prior equivalent to "virtual" counts; often called "smoothing" in NLP and ML.

Mário A. T. Figueiredo (IST & IT)

LxMLS 2013: Probability Theory

# The Bernstein-Von Mises Theorem

• In the previous example, we had n = 10, y = (1, 1, 1, 0, 1, 0, 0, 1, 1, 1), thus  $\sum_i y_i = 7$ . With a Beta prior with  $\alpha = 4$  and  $\beta = 4$ , we had

$$\widehat{x}_{\mathsf{ML}} = 0.7, \quad \widehat{x}_{\mathsf{MAP}} = \frac{3 + \sum_{i} y_{i}}{6 + n} = 0.625, \quad \widehat{x}_{\mathsf{MMSE}} = \frac{4 + \sum_{i} y_{i}}{8 + n} \simeq 0.611$$

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• This illustrates an important result in Bayesian inference: the Bernstein-Von Mises theorem; under (mild) conditions,

$$\lim_{n \to \infty} \widehat{x}_{MAP} = \lim_{n \to \infty} \widehat{x}_{MMSE} = \widehat{x}_{ML}$$

message: if you have a lot of data, priors don't matter.

#### Important Inequalities

• Markov's ineqality: if  $X \ge 0$  is an RV with expectation  $\mathbb{E}(X)$ , then

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• Chebyshev's inequality:  $\mu = \mathbb{E}(Y)$  and  $\sigma^2 = \operatorname{var}(Y)$ , then

$$\mathbb{P}(|X-\mu| \ge s) \le \frac{\sigma^2}{s^2}$$

...simple corollary of Markov's inequality, with  $X = |Y - \mu|^2$ ,  $t = s^2$ 

• Cauchy-Schwartz's inequality for RVs:

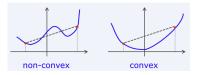
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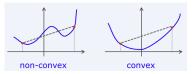


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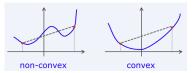
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 $\begin{array}{ll} \text{Examples:} \ \mathbb{E}(X)^2 \leq \mathbb{E}(X^2) \ \Rightarrow \ \text{var}(X) = \mathbb{E}(X^2) - \mathbb{E}(X)^2 \geq 0.\\ \mathbb{E}(\log X) \leq \log \mathbb{E}(X), \ \text{ for } X \text{ a positive RV}. \end{array}$ 

Mário A. T. Figueiredo (IST & IT)

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$$H(X) \le \log K$$
;  
 $H(X) = \log K \Leftrightarrow f_X(x) = 1/k$ , for all  $x \in \{1, ..., K\}$ 

• Measure of uncertainty/randomness of X

Continuous RV X, differential entropy: 
$$h(X) = -\int f_X(x) \log f_X(x) dx$$

• h(X) can be positive or negative. Example, if  $f_X(x) = \text{Uniform}(x; a, b), \ h(X) = \log(b - a).$ 

• If  $f_X(x) = \mathcal{N}(x; \mu, \sigma^2)$ , then  $h(X) = \frac{1}{2}\log(2\pi e\sigma^2)$ .

Entropy of a discrete RV  $X \in \{1, ..., K\}$ :  $H(X) = -\sum_{x=1}^{K} f_X(x) \log f_X(x)$ 

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• If 
$$var(Y) = \sigma^2$$
, then  $h(Y) \leq \frac{1}{2}\log(2\pi e\sigma^2)$ 

Mário A. T. Figueiredo (IST & IT)

Kullback-Leibler divergence (KLD) between two pmf:

$$D(f_X || g_X) = \sum_{x=1}^{K} f_X(x) \log \frac{f_X(x)}{g_X(x)}$$

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Positivity:  $D(f_X || g_X) \ge 0$  $D(f_X || g_X) = 0 \iff f_X(x) = g_X(x)$ , almost everywhere

# Enjoy LxMLS 2013